Scheme for Statistical Analysis of Some Parametric Normalization Classes

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> **Abstract:** In this research 7 parametric classes of normalization functions depending on 1 or 2 parameters proposed for MCDM problem solution. Monte Carlo experiments carried out to perform comparative statistical analysis and find optimal parameter values for the case of Gaussian distribution of decision making matrix elements. Optimal parameter values were ascertained for each normalization method. Normalization formulas were compared with each other in the sense of their efficiency. Logarithmic and Max normalization formulas demonstrated highest values of the best alternative identification. The proposed methodology of determining optimal parameter values of normalization formulas could be applied by approximation of real data with appropriate probability distributions.

> **Keywords:** normalization methods, multi-criteria optimization, Monte Carlo method, comparative statistical analysis, SAW.

1 Introduction

Multiple criteria decision making (MCDM) methods deal with ranking of alternatives (projects) by measurements or evaluations of the projects according to finite number of attributes (criteria). Ranking results depend on many components of this process, the main of them that influence finite results are

- data normalization formula,
- weight determination method,
- data aggregation method.

Data normalization is an important part of a decision-making problem, but it is still not given enough attention in scientific literature. Nevertheless, it was shown in the number of scientific papers that data normalization method selection often significantly affects the accuracy of MCDM problem solution. The main topic of existing articles is investigation of various normalization formulas together with TOPSIS as one of the most popular MCDM method for ranking alternatives. Jahan and Edwards [4] proposed the comprehensive systematized review of existing

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normalization methods. Some of them are traditionally used with the certain MCDM methods, for example, the well known tandem of vector normalization (Van Delft and Nijkamp [17]) and The Technique for Order of Preference by Similarity to Ideal Solution (TOPSIS) method (Hwang and Yoon [3]). The study of Celen [1] revealed that TOPSIS with vector normalization generated the most consistent results among the most popular four normalization procedures. Among the linear normalization procedures, max-min and max methods appeared as the possible alternatives to the vector normalization procedure. A simulation comparison of normalization procedures for TOPSIS in terms of their ranking consistency and weight sensitivity was carried out by Chakraborty and Yeh [2]. The study results also justify the use of the vector normalization procedure for TOPSIS. Influence of normalization tools on COPRAS-G method applied for material selection task proposed by Yazdani *et al.* [19] In the study of Podviezko and Podvezko [13] it is shown that different types of transformation and normalization of data applied to popular MCDA methods, such as SAW or TOPSIS may produce considerable differences in evaluation. Authors stated that attention has to be paid to making a choice of the type of normalization, which reflects preferences of decision-maker.

Kosareva *et al.* [7] accomplished comparative statistical analysis of 5 widely used normalization methods with SAW method for ranking the alternatives and ternary estimates matrix. It is notable that results strongly differ for benefit and cost type attributes. Minmax method in most cases is significantly better than other. In the study of Peldschus [12] the impact of linear, concave and convex function profiles for mapping on a dimensionless interval (normalization) was investigated. Review of the normalization methods used in construction engineering and management, and their applications there are presented by Kaplinski and Tamošaitienė [6]. Milani *et al.* [11] examined how different families of norms affect the result of solving engineering decision problem by entropy and TOPSIS methods. It was verified that the linear optimization norms cannot affect the rank of alternatives significantly. In contrast, nonlinear norms may yield some deviations, mainly for alternatives that are inherently close.

Recently, some new normalization formulas were proposed in the literature. Research of Zavadskas and Turskis [20] is focused on introducing a new logarithmic method for decision making matrix normalization. Based on Weitendorf [18] and Juttler [5] formulas, Stanujkic *et al.* [14] proposed a new normalization procedure. The idea is to use the distance from the preferred ratings, which respects the decision-maker's preferences. This procedure, adapted for negotiations, was integrated with Step-Wise Weight Assessment Ratio Analysis (SWARA) and Additive Ratio Assessment (ARAS) methods in the research of Stanujkic *et al.* [15] Data normalization, as well as the measurement scales, inconsistency issues, missing judgement estimation methods, etc. have been extensively studied in the pairwise comparison matrix (PCM). A comprehensive literature review on PCM provided in Kou *et al.* [8] A group decision-making (GDM) method for integrating heterogeneous information proposed by Li *et al.* [9]

The purpose of this study is to propose a new methodology of constructing parametric normalization methods and to carry out their statistical comparative analysis. 7 classes of parametric data normalization procedures are presented. Some of the proposed normalization methods with particular parameter values are well known and widespread, other methods aren't so popular. Nevertheless, all these methods were not being applied using the wide range of parameter values so far.

The article is organized as follows. In the Section 2 seven classes of parametric normalization functions are introduced, their properties and dependency on parameter values are discussed. Experiment design and detailed description of initial data matrices generation procedure is given in the Section 3. Monte Carlo experiment results and comparative statistical analysis of normalization methods depending on the parameter values presented in the Section 4. Conclusions and future research are discussed in the Section 5.

2 Parametric data normalization methods

In this article Simple Additive Weighting (SAW) method (see Ref. 19) with equal weights is applied for MCDM problem solution. Let us suppose that the initial data are the results of some measurements, expert evaluations, etc., and are written in $m \times n$ -dimension matrix $X = (x_{ij})_{m \times n}$. The element x_{ij} of decision making matrix is evaluation of alternative i(i = 1, 2, ..., m) by the criterion j(j = 1, 2, ..., n). Decision making matrix after normalization procedure is noted as follows: $\tilde{X} = (\tilde{x}_{ij})_{m \times n}, 0 \leq \tilde{x}_{ij} \leq 1$. Let $w_j, j = 1, 2, ..., n$ be criteria weights satisfying conditions $\sum_{j=1}^{n} w_j, 0 \leq w_j \leq 1$. Then, SAW criteria aggregated value is calculated for each alternative:

$$Q_i = \sum_{j=1}^n w_j \tilde{x}_{ij}, i = 1, 2, \dots, m.$$

If $Q_1 \ge Q_2 \ge \ldots \ge Q_m$, then the alternatives are ranked as follows:

$$altern_1 \succ altern_2 \succ \ldots \succ altern_m$$

The essence of any normalization procedure is mapping of the real values $x_{ij} \in [m_j, M_j] \subset (-\infty, +\infty)$ having certain meaning in the interval [0, 1], which in general case represents these numbers unevenly. For example, functions $x^{\alpha} : [0, 1] \rightarrow [0, 1]$ depending on the parameter α represent half of the maximum values to the 13% of the maximum normalized values, when $\alpha = 0.2$, to the 30%, when $\alpha = 0.5$, 75%, when $\alpha = 2$ and even 97%, when $\alpha = 5$ (see Figure 1).



Figure 1: Normalization functions $x^{\alpha} : [0, 1] \to [0, 1]$.

It means that the aggregated values Q_i can not only depend strongly and not so much on the absolute values of x_{ij} as on the relationships $x_{i'j} < x_{i''j}$. In this research we'll limit ourselves with direct optimization, when higher values of criteria are better, omitting inverse optimization, when lower values are treated as better. So, the only essential requirement for normalization functions is – they must be monotonously increasing. The other restriction of our research is dealing with random generation of initial decision making matrix elements that have the Gaussian (normal) probability distribution.

Practically applied normalization methods can be more complex compared to functions depicted in Figure 1 and can therefore have a great deal of influence in alternative comparisons and decision-making results. Therefore, the literature deals with a large number of normalization methods, the new methods are developed, their researches and comparisons are carried out.

Denote $m_j = \min_{1 \le i \le m} x_{ij}, M_j = \max_{1 \le i \le m} x_{ij}, j = 1, 2, ..., n$. In the Table 1 we propose 7 classes of normalization functions $[m_j, M_j] \to [0, 1]$ depending on 2 parameters $\alpha \ge 0$ and $k \ge 0$.

These classes describe 7 different normalization approaches. Note, that (1), (6) and (7) methods depend exclusively on the parameter α , meanwhile the rest (2)–(5) methods are depending on two parameters $\alpha \geq 0$ and $k \geq 0$. When $\alpha = 1$, methods (1), (6) and (7) are known from the literature and are entitled as: (1) – Minmax normalization, (6) – Logaritmic normalization, (7) – Max normalization. Plots of functions (1)–(7), when $\alpha = 1, k = 1$, are given in Figure 2, plots of functions (1), (6) and (7), when $\alpha = 0.5, 2$ – in Figure 3). Graphs of functions (2)–(5) with corresponding parameter values $\alpha = 0.5, 2; k = 2, 3, 5$ are depicted in Figure 4. In the Table 2 the example of data matrix normalization by methods (1)–(7) is presented when $\alpha = 1, k = 1$ and initial data matrix is

$$(x_{ij})_{(4\times4)} = \begin{pmatrix} 120 & 1.2 & 1 & 4\\ 250 & 2.5 & 2 & 5\\ 3600 & 36 & 5 & 8\\ 6400 & 64 & 10 & 13 \end{pmatrix}$$

The second column elements of matrix (x_{ij}) are obtained dividing the first column elements by 100, while the fourth column elements are calculated by adding constant 3 to the third column elements. Let us see that data normalization formulas (1)–(5) are invariant with respect to linear transformation $\alpha x + \beta$, formula (7) is invariant with respect to multiplication/division, but does not preserve addition, while formula (6) does not preserve nor multiplication, neither addition. All the first 5 methods map the lowest value to 0, and the highest value to 1, method (7) maps the highest value to 1.

Several questions naturally arise considering this issue. Are the mutual differences between the results of these methods application for MCDM problem solution significant? What are the "best" values of parameters α and k? How do the results vary when varying parameter α and kvalues? How do the results vary when varying m and n values?

| Table 1 | : | Formul | as | for | 7 | parametric | normal | izatior | 1 methods | in | the | case o | f (| direct | norma | lizat | ion. |
|---------|---|--------|----|-----|---|------------|--------|---------|-----------|----|-----|--------|-----|--------|-------|-------|------|
|---------|---|--------|----|-----|---|------------|--------|---------|-----------|----|-----|--------|-----|--------|-------|-------|------|

| Formula number | Normalization method | Direct normalization formula |
|-------------------|---|--|
| (1) | Minmax normalization (Weitendorf, 1976) | $\tilde{x}_{ij} = \left(\frac{x_{ij} - m_j}{M_j - m_j}\right)^{\alpha}$ |
| | | $-k\left(rac{M_j-x_{ij}}{m_j} ight)^{lpha}$ |
| (2) | Exponential normaliza- tion | $\tilde{x}_{ij} = e \left(\begin{array}{c} x_{ij} - m_j \end{array} \right)$ |
| (3) | Logaritmic maxmin nor- malization | $\tilde{x}_{ij} = \frac{1}{\left(1 + k \ln\left(\frac{M_j - m_j}{r_{ij} - m_j}\right)\right)^{\alpha}}$ |
| (4) | Arctangent normaliza- tion | $\tilde{x}_{ij} = \left(\frac{2}{\pi}\arctan\left(k\frac{x_{ij}-m_j}{M_j-x_{ij}}\right)\right)^{\alpha}$ |
| (5) | Double exponential nor- malization | $\tilde{x}_{ij} = \left(\frac{1-e^{-k\left(\frac{x_{ij}-m_j}{M_j-x_{ij}}\right)}}{\frac{1-e^{-k\left(\frac{x_{ij}-m_j}{M_j-x_{ij}}\right)}}\right)^{\alpha}$ |
| (6) | Logarithmic normal- ization (Zavadskas and Turskis, 2008) | $\tilde{x}_{ij} = \left(\frac{\ln(x_{ij})}{\ln(\prod_{i=1}^m x_{ij})}\right)^{\alpha}, x_{ij} \ge 1$ |
| (7) | Max normalization (Stopp [16], 1975) | $\tilde{x}_{ij} = \left(\frac{x_{ij}}{M_j}\right)^{\alpha}$ |



Figure 2: Functions (1)–(7), when $\alpha = 1, k = 1$.



Figure 3: Functions (1), (6) and (7), when $\alpha = 0.5$ and $\alpha = 2$.

3 Experiment design

The case under investigation is when the decision making matrix elements are some measurements done with sufficient precision and having Gaussian probability distribution $x_{ij} \sim N(\mu_i, \sigma)$. We'll suppose, that the first row elements of the matrix have at the average higher values than elements of other rows $\mu_1 > \mu_2 = \ldots = \mu_m$. Only such matrices were analysed, for which the first alternative (first row) has no domination property in comparison with any other alternative, i.e. when the following conditions are not fulfilled:

$$(x_{11}, x_{12}, \dots, x_{1n}) \succ (x_{i1}, x_{i2}, \dots, x_{in}), \text{ if } x_{11} \ge x_{i1}, x_{12} \ge x_{i2}, \dots, x_{1n} \ge x_{in}.$$
 (8)

If condition (8) is valid, weighted averages $Q_i = \sum_{j=1}^n w_j \tilde{x}_{ij}$ will satisfy inequalities $Q_1 \ge Q_i$ with any values of weights w_j . Therefore, the result of all normalization methods will be the same – the first alternative is better than *i*-th alternative. So, when random matrices are being generated, such matrices whose first and any (at least one) *i*-th row satisfy domination property (8) are being rejected. As a result, the number of cases of a fair decision has decreased. For example, the first and the second rows elements of decision making matrix

$$X^{(1)} = \begin{pmatrix} 52.34 & 66.31 & 63.38 & 67.01 \\ 48.90 & 62.05 & 56.54 & 53.93 \\ 72.05 & 56.24 & 61.58 & 48.05 \\ 52.10 & 65.00 & 71.95 & 56.82 \\ 77.69 & 70.34 & 65.00 & 55.65 \end{pmatrix}$$

| Normalization method | Normalized matrix \tilde{x}_{ij} | Formula |
|-------------------------|--|--|
| (1) | $\begin{pmatrix} 0 & 0 & 0 & 0 \\ 0.021 & 0.021 & 0.11 & 0.11 \\ 0.55 & 0.55 & 0.44 & 0.44 \\ 1 & 1 & 1 & 1 \end{pmatrix}$ | $\tilde{x}_{ij} = \frac{x_{ij} - m_j}{M_j - m_j}$ |
| (2) | $\begin{pmatrix} 0 & 0 & 0 & 0 \\ 2.8 \cdot 10^{-21} & 2.8 \cdot 10^{-21} & 0.0003 & 0.0003 \\ 0.45 & 0.45 & 0.29 & 0.29 \\ 1 & 1 & 1 & 1 \end{pmatrix}$ | $\tilde{x}_{ij} = e^{-\left(\frac{M_j - x_{ij}}{x_{ij} - m_j}\right)}$ |
| (3) | $\begin{pmatrix} 0 & 0 & 0 & 0 \\ 0.21 & 0.21 & 0.31 & 0.31 \\ 0.63 & 0.63 & 0.55 & 0.55 \\ 1 & 1 & 1 & 1 \end{pmatrix}$ | $\tilde{x}_{ij} = \frac{1}{1 + \ln\left(\frac{M_j - m_j}{x_{ij} - m_j}\right)}$ |
| (4) | $\begin{pmatrix} 0 & 0 & 0 & 0 \\ 0.01 & 0.01 & 0.08 & 0.08 \\ 0.57 & 0.57 & 0.43 & 0.43 \\ 1 & 1 & 1 & 1 \end{pmatrix}$ | $ \tilde{x}_{ij} = \frac{2}{\pi} \arctan\left(\frac{x_{ij} - m_j}{M_j - x_{ij}}\right) $ |
| (5) | $\begin{pmatrix} 0 & 0 & 0 & 0 \\ 0.01 & 0.01 & 0.06 & 0.06 \\ 0.55 & 0.55 & 0.38 & 0.38 \\ 1 & 1 & 1 & 1 \end{pmatrix}$ | $\tilde{x}_{ij} = \frac{1 - e^{-\left(\frac{x_{ij} - m_j}{M_j - x_{ij}}\right)}}{1 + e^{-\left(\frac{x_{ij} - m_j}{M_j - x_{ij}}\right)}}$ |
| (6) | $\begin{pmatrix} 0.18 & 0.02 & 0 & 0.18 \\ 0.20 & 0.10 & 0.15 & 0.21 \\ 0.30 & 0.41 & 0.35 & 0.27 \\ 0.32 & 0.47 & 0.5 & 0.34 \end{pmatrix}$ | $\tilde{x}_{ij} = \frac{\ln(x_{ij})}{\ln(\prod_{i=1}^m x_{ij})}$ |
| (7) | $\begin{pmatrix} 0.02 & 0.02 & 0.1 & 0.31 \\ 0.04 & 0.04 & 0.2 & 0.38 \\ 0.56 & 0.56 & 0.5 & 0.62 \\ 1 & 1 & 1 & 1 \end{pmatrix}$ | $\tilde{x}_{ij} = \frac{x_{ij}}{M_j}$ |

Table 2: Normalization methods application example when $\alpha = 1$ and k = 1.

satisfy inequalities $52.34 \ge 48.90, 66.31 \ge 62.05, 63.38 \ge 56.54, 67.01 \ge 53.93$,

therefore $(52.34, 66.31, 63.38, 67.01) \succ (48.90, 62.05, 56.54, 53.93)$ and matrix $X^{(1)}$ is being rejected. Matrix

$$X^{(2)} = \begin{pmatrix} 57.56 & 66.42 & 58.31 & 55.51 \\ 58.51 & 63.08 & 67.93 & 46.31 \\ 55.36 & 57.58 & 59.16 & 72.82 \\ 38.95 & 58.40 & 40.93 & 57.93 \\ 53.68 & 45.57 & 70.41 & 59.98 \end{pmatrix}$$

is appropriate. The results of 7 normalization procedures when $\alpha = 1, k = 1$ are as follows:

$$\tilde{X}^{(2)}(1) = \begin{pmatrix} 0.95 & 1.0 & 0.59 & 0.35\\ 1.0 & 0.84 & 0.92 & 0.0\\ 0.84 & 0.58 & 0.62 & 1.0\\ 0.0 & 0.62 & 0.0 & 0.44\\ 0.75 & 0.0 & 1.0 & 0.52 \end{pmatrix}, \quad \tilde{X}^{(2)}(2) = \begin{pmatrix} 0.95 & 1.0 & 0.50 & 0.15\\ 1.0 & 0.83 & 0.91 & 0.0\\ 0.83 & 0.48 & 0.54 & 1.0\\ 0.0 & 0.54 & 0.0 & 0.28\\ 0.72 & 0.0 & 1.0 & 0.39 \end{pmatrix},$$



Figure 4: Graphs of functions (2)–(5) with different parameter α and k values.

$$\begin{split} \tilde{X}^{(2)}(3) &= \begin{pmatrix} 0.95 & 1.0 & 0.65 & 0.49 \\ 1.0 & 0.85 & 0.91 & 0.0 \\ 0.85 & 0.65 & 0.68 & 1.0 \\ 0.0 & 0.67 & 0.0 & 0.55 \\ 0.78 & 0.0 & 1.0 & 0.60 \end{pmatrix}, \quad \tilde{X}^{(2)}(4) = \begin{pmatrix} 0.97 & 1.0 & 0.61 & 0.31 \\ 1.0 & 0.88 & 0.94 & 0.0 \\ 0.88 & 0.60 & 0.65 & 1.0 \\ 0.0 & 0.64 & 0.0 & 0.42 \\ 0.80 & 0.0 & 1.0 & 0.52 \end{pmatrix}, \\ \tilde{X}^{(2)}(5) &= \begin{pmatrix} 1.0 & 1.0 & 0.62 & 0.26 \\ 1.0 & 0.99 & 1.0 & 0.0 \\ 0.99 & 0.59 & 0.67 & 1.0 \\ 0.0 & 0.66 & 0.0 & 0.37 \\ 0.91 & 0.0 & 1.0 & 0.49 \end{pmatrix}, \quad \tilde{X}^{(2)}(6) = \begin{pmatrix} 0.21 & 0.21 & 0.20 & 0.20 \\ 0.21 & 0.20 & 0.21 & 0.19 \\ 0.20 & 0.20 & 0.20 & 0.21 \\ 0.19 & 0.20 & 0.18 & 0.20 \\ 0.20 & 0.19 & 0.21 & 0.20 \end{pmatrix}, \\ \tilde{X}^{(2)}(7) &= \begin{pmatrix} 0.98 & 1.0 & 0.83 & 0.76 \\ 1.0 & 0.95 & 0.96 & 0.64 \\ 0.95 & 0.87 & 0.84 & 1.0 \\ 0.67 & 0.88 & 0.58 & 0.80 \\ 0.92 & 0.69 & 1.0 & 0.82 \end{pmatrix}. \end{split}$$

We can see that matrices $\tilde{X}^{(2)}(1) - \tilde{X}^{(2)}(5)$ mutually differ slightly, meanwhile there are essential differences between the first 5 matrices and $\tilde{X}^{(2)}(6) - \tilde{X}^{(2)}(7)$. When proper matrix X is generated and some kind of normalization (1)–(7) is done, SAW criteria aggregated values with

equal weights are calculated for each row:

$$Q_i = \sum_{j=1}^n \frac{1}{n} \cdot \tilde{x}_{ij}, \, i = 1, 2, \dots, m.$$
(9)

The first alternative is considered to be the best one when $Q_1 > Q_2$, $Q_1 > Q_3, \ldots, Q_1 > Q_m$. 10 series by 100 experiments were conducted. After 100 realizations of experiments denote ID the number of cases, when the first alternative identified as the best one (the number of correct MCDM problem solutions from 100). Denote also $W_{ID} = \frac{ID}{100}$, i.e. the percent of the correct MCDM problem solutions.

Table 3: Numbers of cases with the best first alternative (ID) for normalization methods (1)–(7), when $\alpha = 1$ and k = 1.

| (1) | 62 | 54 | 62 | 56 | 58 | 52 | 49 | 58 | 52 | 45 |
|-----|----|----|----|----|----|----|----|----|----|----|
| (2) | 69 | 60 | 66 | 57 | 56 | 56 | 50 | 57 | 54 | 46 |
| (3) | 55 | 50 | 59 | 55 | 51 | 45 | 45 | 57 | 55 | 44 |
| (4) | 64 | 56 | 60 | 58 | 56 | 51 | 50 | 57 | 52 | 46 |
| (5) | 62 | 56 | 59 | 59 | 53 | 49 | 47 | 53 | 51 | 47 |
| (6) | 69 | 66 | 66 | 73 | 63 | 64 | 58 | 64 | 63 | 56 |
| (7) | 69 | 65 | 66 | 72 | 63 | 62 | 61 | 65 | 63 | 54 |

Table 3 shows experiment results – how many times after application of methods (1)–(7) with $\alpha = 1$ and k = 1 it was found that the best option is the first alternative. Each row contains results of 10 experiment series for each normalization method.

Next, Monte Carlo experiments were conducted by changing parameters α and k values as follows. First row elements of matrix X were generated according to the Gaussian distribution with the average value $\mu_1 = 67$ and standard deviation $\sigma = 15$, elements of other rows – with the average values $\mu_i = 57$, i = 2, 3, ..., m and standard deviation $\sigma = 15$. Calculations were performed using the C++ program.

4 Experiment results

4.1 Dependence of the best alternative detection accuracy on parameter α

At the first stage of the experiments we fixed k value at k = 1 and varied α choosing the individual range for each normalization method. 100 series of 100 experiments were repeated for each parameter α value and empirical averages $\overline{W}_{ID} = \frac{1}{100} \sum_{i=1}^{100} W_{ID_i}$ calculated. The purpose is to detect α value which maximizes \overline{W}_{ID} .

In the Table 4 the dynamics of \overline{W}_{ID} values depending on α is depicted. 95% confidence intervals $[w_{1_0.95}; w_{2_0.95}]$ for mean values EW_{ID} are given in the last column of the table. Maximum values of correct MCDM problem solution empirical averages \overline{W}_{ID} were achieved for such α values: $\alpha = 2$ for methods (1), (3), (4) and (7), $\alpha = 0.75$ for method (2), $\alpha = 4$ for method (5), $\alpha = 20$ for method (6). The dependence of empirical mean values on α as well as upper and lower bounds of confidence intervals of the mean values are presented graphically in Figs. 5–6. For improvement of graphic images, smoothing spline fitting of Table 4 data was produced with MATLAB.

The question arises: are the detected differences between normalization methods with different α values statistically significant? Student's t-tests were applied for testing the hypotheses H_0 of equal average values EW_{ID} in the cases of $\alpha = 1$ and corresponding optimal parameter α values. Table 5 shows *p*-values of the respective *t*-tests. Significant differences in the best alternative

| Normalization method | α | \overline{W}_{ID} | $[w_{1_0.95}; w_{2_0.95}]$ |
|-------------------------|--------|---------------------|---|
| (1) | 1 | 0.4784 | [0.4678; 0.4890] |
| | 1.25 | 0.4899 | [0.4790; 0.5008] |
| | 1.5 | 0.4998 | [0.4885; 0.5111] |
| | 2. | 0.5015 | [0.4911; 0.5119] |
| | 3 | 0.4958 | [0.4860; 05055] |
| | 4 | $0,\!4905$ | [0.4804; 0.5006] |
| | 5 | 0.4859 | [0.4758; 0.4960] |
| (2) | 0.5 | 0.4952 | [0.4854; 0.5050] |
| | 0.75 | 0.4987 | [0.4876; 0.5098] |
| | 1 | 0.4936 | [0.4836; 0.5036] |
| | 1.25 | 0.4897 | [0.4796; 0.4998] |
| | 1.5 | 0.4874 | [0.4774; 0.4974] |
| | 2 | 0.4807 | [0.4707; 0.4907] |
| (3) | 0.5 | 0.3438 | [0.3342; 0.3534] |
| | 1 | 0.4485 | [0.4373; 0.3534] |
| | 1.5 | 0.4859 | [0.4758; 0.4597] |
| | 2 | 0.4979 | [0.4875; 0.4960] |
| | 3 | 0.4959 | [0.4862; 0.5083] |
| | 4 | 0.4919 | [0.4821: 0.5017] |
| (4) | 1 | 0.4751 | $\frac{[0.4643: 0.4859]}{[0.4643: 0.4859]}$ |
| (-) | 1.5 | 0.4947 | [0.4842; 0.5052] |
| | 2 | 0.5009 | [0.4904: 0.5114] |
| | 3 | 0.498 | [0.4880; 0.5080] |
| | 4 | 0.4929 | [0.4827; 0.5031] |
| | 5 | 0.4906 | [0.4803; 0.5009] |
| (5) | 1 | 0.4684 | [0.4582: 0.4786] |
| (*) | 2 | 0.4914 | [0.4812; 0.5016] |
| | 3 | 0.4951 | [0.4845; 0.5057] |
| | 4 | 0.4952 | [0.4844: 0.5060] |
| | 5 | 0.494 | [0.4834: 0.5046] |
| | 6 6 | 0.4927 | [0.4823; 0.5031] |
| | 7 | 0 4923 | [0.4822; 0.5024] |
| (6) | 1 | 0.5370 | $\frac{[0.1022, 0.0021]}{[0.5258; 0.5482]}$ |
| (0) | 2 | 0.5396 | [0.5287; 0.5505] |
| | - 3 | 0.5355 0.5444 | [0.5336: 0.5552] |
| | 4 | 0.5486 | [0.5377: 0.5595] |
| | 5 | 0.5508 | [0.5401; 0.5616] |
| | 10 | 0.5506 | [0.5496: 0.5696] |
| | 20 | 0.5609 | [0.5512: 0.5706] |
| | 30 | 0.5523 | [0.5012, 0.5700] |
| (7) | 0.25 | 0.5304 | [0.5284. 0.55040] |
| (\mathbf{r}) | 0.20 | 0.5594 | [0.5264, 0.5504] [0.5200· 0.5510] |
| | 0.5 | 0.5400 | [0.5250, 0.5510] [0.5286: 0.5506] |
| | 1 | 0.0090 | [0.5260, 0.5500] [0.5203: 0.5512] |
| | ⊥ ງ | 0.0400 | [0.0290; 0.0010] [0.5310: 0.5599] |
| | ⊿ ? | 0.5419 | [0.0010; 0.0028] [0.5076: 0.5400] |
| | ა 4 | 0.0004 | [0.5270; 0.5492] [0.5220; 0.5449] |
| | 4 | 0.3340 | [0.3232; 0.3448] |

Table 4: \overline{W}_{ID} values and confidence intervals for mean values EW_{ID} for 7 normalization methods at k = 1.

detection accuracy on significance level 0.05 were observed for normalization methods (1), (3)–(6). Methods (2) and (7) didn't show significant differences. Consequently, it makes sense using formulas (1), (3)–(6) with the appropriate optimal parameter α values that differ from $\alpha = 1$ (see optimal α values in the Table 5), while it makes sense using $\alpha = 1$ for methods (2) and (7).



Figure 5: Dependence of empirical average accuracy \overline{W}_{ID} on the parameter α values for (1)–(6) normalization methods and 95% confidence intervals bounds.

Next, ANOVA procedure was applied to check the hypothesis of equal averages EW_{ID} for 7 normalization methods with optimal α values. Procedure results show significant differences existing between methods. Bonferroni test was chosen for Post Hoc multiple comparisons. It revealed that there are nor significant differences between averages EW_{ID} for (1)–(5) methods, neither between (6) and (7) methods. However, when comparing any method from the first group



Figure 6: Dependence of empirical average Figure 7: Mean plot from ANOVA proaccuracy \overline{W}_{ID} on the parameter α values cedure comparing EW_{ID} for optimal α for (7) method and 95% confidence inter-values corresponding to 7 normalization vals bounds.

methods.

((1)-(5) methods) with any of the methods of the second group ((6)-(7) methods), statistically significant differences were found. The conclusion can be made that (6) and (7) methods at the average are more precise than (1)-(5) methods, since the average accuracy EW_{ID} obtained by (6)-(7) methods is significantly higher than the accuracy of (1)-(5) normalization formulas. In Figure 7 the mean plot from ANOVA procedure output is represented. EW_{ID} for optimal α values corresponding to 7 normalization methods are depicted.

Table 5: Hypotheses of equal average values EW_{ID} testing results for $\alpha = 1$ and optimal α value for (1)-(7) normalization methods for the first experiment.

| Normalization | Optimal | Ho | <i>n</i> -value |
|---------------|----------------|--|-----------------|
| method | α value | | p varae |
| (1) | 2 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=2}$ | 0.002 |
| (2) | 0.75 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=0.75}$ | 0.499 |
| (3) | 2 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=2}$ | 0.000 |
| (4) | 2 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=2}$ | 0.001 |
| (5) | 4 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=4}$ | 0.000 |
| (6) | 20 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=20}$ | 0.002 |
| (7) | 2 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=2}$ | 0.838 |

Table 6: \overline{W}_{ID} values, corresponding optimal α values and t-test results for (1)–(7) normalization methods for the second experiment.

| Normalizatio | onOptima | $1 \overline{W}_{ID}$ | H_0 | <i>p</i> -value |
|--------------|----------|-----------------------|--|-----------------|
| method | α | | | |
| | value | | | |
| (1) | 2 | 0.6967 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=2}$ | 0.000 |
| (2) | 0.75 | 0.6966 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=0.75}$ | 0.594 |
| (3) | 2 | 0.6959 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=2}$ | 0.000 |
| (4) | 2 | 0.701 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=2}$ | 0.000 |
| (5) | 3 | 0.6879 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=3}$ | 0.002 |
| (6) | 10 | 0.7775 | $EW_{ID \ \alpha=1} = EW_{ID \ \alpha=20}$ | 0.003 |
| (7) | 1 | 0.7631 | - | - |

Table 7: Recommended normalization formulas (1)–(7) with optimal α values and k = 1.

$$\tilde{x}_{ij} = \left(\frac{x_{ij} - m_j}{M_j - m_j}\right)^2 \tag{1}$$

$$\tilde{x}_{ij} = e^{-\left(\frac{M_j - x_{ij}}{x_{ij} - m_j}\right)} \tag{2}$$

$$\tilde{x}_{ij} = \frac{1}{\left(1 + \ln\left(\frac{M_j - m_j}{x_{ij} - m_j}\right)\right)^2} \tag{3}$$

$$\tilde{x}_{ij} = \left(\frac{2}{\pi} \arctan\left(\frac{x_{ij} - m_j}{M_j - x_{ij}}\right)\right)^2 \qquad (4)$$

$$\tilde{x}_{ij} = \begin{pmatrix} \frac{-\binom{x_{ij} - m_j}{M_j - x_{ij}}}{1 + e^{-\binom{x_{ij} - m_j}{M_j - x_{ij}}}} \end{pmatrix}$$
(5)

$$\tilde{x}_{ij} = \left(\frac{\ln(x_{ij})}{\ln(\prod_{i=1}^m x_{ij})}\right)^{20}, x_{ij} \ge 1 \qquad (6)$$
$$\tilde{x}_{ij} = \frac{x_{ij}}{M_j} \qquad (7)$$

Due to randomness of the experiments we observe some variability in the obtained results: optimal α value for (5)-th normalization method changed from 4 to 3, for (6)-th method – from 20 to 10, for (7)-th method – from 2 to 1. However, *t*-test didn't detect significant differences between EW_{ID} at the mentioned α levels. ANOVA results and conclusions also are the same. So, the conclusions remained unchained – it is recommended to use optimal α values, which differ from 1 for normalization methods (1), (3)–(6) and apply $\alpha = 1$ for methods (2) and (7). Moreover, (6) and (7) normalization formulas lead to significantly higher average accuracy of best alternative detection. The results of the experiments revealed that it is reasonable to use normalization formulas with α values specified in the Table 7.

To evaluate the stability of obtained results, Monte Carlo experiments were repeated with other values of decision making matrices. The elements of matrix X were generated as follows: $x_{1j} \sim N(67, 15), x_{ij} \sim N(52, 15), i = 2, 3, \ldots, m$. As the differences $\mu_1 - \mu_i$ increased, it is natural to expect that the best alternative detection accuracy will be higher than in the previous experiment. The results of the second experiment are shown in the Table 6. They essentially confirmed results of the first experiment.

4.2 Dependence of the best alternative detection accuracy on parameter k

Data normalization formulas (2)–(5) are depending on both parameters α and k. So, it is interesting to reveal whether EW_{ID} change significantly while changing k. Calculations were carried out with few values of parameters α , and for each fixed α for few values of k. The elements of matrix X were generated as follows: $x_{1j} \sim N(67, 15), x_{ij} \sim N(57, 15), i = 2, 3, \ldots, m$. 100 Monte Carlo experiments were repeated for each parameter values combination and empirical mean values \overline{W}_{ID} calculated. Then maximum value of \overline{W}_{ID} was detected for some optimal parameter k value. Calculation results for normalization formulas (4) and (5) are presented in the Table 8. Maximum \overline{W}_{ID} values and corresponding parameter α and k values are as follows: normalization method (2) – $\overline{W}_{ID} = 0.511, \alpha = 0.75, k = 1$; (3) – $\overline{W}_{ID} = 0.505, \alpha = 2, k = 1$; (4) – $\overline{W}_{ID} = 0.516, \alpha = 2, k = 1$; (5) – $\overline{W}_{ID} = 0.509, \alpha = 1, k = 0.5$. In Figure 8 the dependence of average accuracy \overline{W}_{ID} on k = 1 is depicted. For normalization formulas (2)–(4) optimal parameter k value is 1, only for method (5) optimal value differs from 1 (k = 0.5).

Next, t-test was applied for testing the hypotheses H_0 of equal average values EW_{ID} for k = 1 and corresponding optimal parameter k values. The difference between EW_{ID} with parameters $\alpha = 1, k = 0.5$ and EW_{ID} with parameters $\alpha = 1, k = 1$ is significant at significance level 0.05.



Figure 8: Dependence of empirical average accuracy \overline{W}_{ID} on parameter k values for (2)–(5) normalization methods.

| Method (4) | | | | | | | | | |
|------------|-------|-------|------------|-------|----------|-------|--|--|--|
| α | 0.5 | 0.75 | 1 | 1.5 | 2 | 3 | | | |
| k | | | | | | | | | |
| 0.1 | 0.497 | 0.486 | 0.500 | 0.482 | 0.486 | 0.479 | | | |
| 0.2 | 0.510 | 0.494 | 0.500 | 0.496 | 0.484 | 0.479 | | | |
| 0.5 | 0.475 | 0.500 | 0.501 | 0.509 | 0.495 | 0.497 | | | |
| 1 | 0.422 | 0.460 | 0.486 | 0.502 | 0.516 | 0.506 | | | |
| 2 | 0.353 | 0.399 | 0.435 | 0.478 | 0.488 | 0.504 | | | |
| 5 | 0.303 | 0.317 | 0.358 | 0.381 | 0.401 | 0.449 | | | |
| 10 | 0.285 | 0.297 | 0.307 | 0.324 | 0.345 | 0.378 | | | |
| | | Met | thod (5) | | | | | | |
| α | 0.5 | 0.75 | 1 | 1.5 | 2 | 3 | | | |
| k | | | | | | | | | |
| 0.1 | 0.487 | 0.482 | 0.479 | 0.485 | 0.484 | 0.485 | | | |
| 0.2 | 0.497 | 0.483 | 0.479 | 0.486 | 0.484 | 0.483 | | | |
| 0.5 | 0.483 | 0.502 | 0.509 | 0.504 | 0.497 | 0.498 | | | |
| 1 | 0.423 | 0.466 | 0.482 | 0.489 | 0.494 | 0.494 | | | |
| 2 | 0.339 | 0.385 | 0.411 | 0.442 | 0.454 | 0.473 | | | |
| 5 | 0.296 | 0.302 | 0.318 | 0.339 | 0.354 | 0.381 | | | |
| 10 | 0.287 | 0.278 | 0.290 | 0.299 | 0.296 | 0.323 | | | |

Table 8: \overline{W}_{ID} calculated by varying α and k values for normalization methods (4) and (5).

So, it is recommended to use $\alpha = 1, k = 0.5$ for normalization method (5):

$$\tilde{x}_{ij} = \frac{1 - e^{-\frac{1}{2} \left(\frac{x_{ij} - m_j}{M_j - x_{ij}}\right)}}{1 + e^{-\frac{1}{2} \left(\frac{x_{ij} - m_j}{M_j - x_{ij}}\right)}},$$

whereas for methods (2)–(4) it is reasonable to use k = 1 with respective optimal α value (see relevant formulas in Table 7).

5 Conclusions and future research

The purpose of this article is to analyze some parametric normalization formulas and establish how various data normalization methods and parameter values affect the accuracy of MCDM problem solution. 7 data normalization methods were investigated (see formulas (1)-(7)). Some of them with certain parameter values are generalizations of the well known normalization methods. Data matrices were randomly generated according to Gaussian probability distribution. In all conducted Monte Carlo experiments decision making matrices were generated with the first alternative as the best one. Then the alternatives were ranked by the SAW method overall aggregated value with equal weights. The measure of identification accuracy is the percentage of correct identifications of the best alternative. The results of experiments are as follows.

- 1. Identification accuracy obtained by methods (6)–(7) is significantly higher than the accuracy for normalization methods (1)–(5).
- 2. Variation of parameter k revealed that for normalization method (5) it is reasonable to use combination of parameters $\alpha = 1, k = 0.5$, whereas for methods (2)–(4) it is reasonable to use k = 1 with respective optimal α value (see relevant formulas in Table 7).
- 3. Optimal α value have some "degrees of freedom", for example, it is possible to choose another optimal value $\alpha = 10$ instead of $\alpha = 20$ for normalization formula (6), since there isn't significant difference between identification accuracy at these values.
- 4. This research is accomplished in the special case when elements of decision matrix have Gaussian distribution. If we possess more information, we can apply the same methodology by approximating real data with appropriate probability distributions.
- 5. Parameters of initial decision making matrix were chosen so that identification of the best alternative would be higher than 50%. Experiments revealed that identification accuracy is higher with the bigger number of criteria n, and is lower with fewer alternatives m.

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